### **OIKONOMIKO** ΠΑΝΕΠΙΣΤΗΜΙΟ **AOHNON**



ATHENS UNIVERSITY OF ECONOMICS **AND BUSINESS** 

### **Department of Statistics**

### **ECONOMETRICS Panel Data Models**

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# **Panel Data Models**

- Introduction and motivation
- Panel data models
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	- Random effects panel model
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### **Introduction and motivation**

Advantages:

- Larger number of data points (more degrees of freedom)
- Combine cross section and time series data
- Limitation of the omitted variable problem
- Takes into account common and cross sectional explanatory variables
- May account for cross sectional dependence

# **Panel data**

### **GDP series for the G7 countries**



## **Panel data**

**Dependent variables and Specific cross sectional explanatory variables**

**Panel data:** 

**Dependent variable**:  $Y_{it}$ ,  $i = 1, ..., N$  (cross section unit  $t = 1, ..., T$  (time period)

**Explanatory variables** (specific cross sectional unit variables, i.e. different for each unit)

 $X_{ijt}$ ,  $i = 1, ..., N$  (cross section unit  $j = 1, \ldots, k$  (explanatory variables)  $t = 1, ..., T$  (time period)

E.g.:  $Y_{it}$ : GDP for different countries across time,  $X_{ijt}$ : investments, industrial production, unemployment for different countries across time





### **Dependent variables and Common explanatory variables**

**Panel data:** 

**Dependent variable:**  $Y_{it}$ ,  $i = 1, ..., N$  (cross section unit  $t = 1, ..., T$  (time period)

 **Explanatory variables** (common variables, i.e. common for each unit)

 $X_{jt}$  $X_{it}$ ,  $j = 1, ..., k$  (explanatory variables)  $t = 1, ..., T$  (time period)

E.g.:  $Y_{it}$ : GDP for different countries across time,  $X_{it}$ : global volatility index, global panic index across time (common explanatory variables for all units)



### **A general dynamic panel data model**

Consider the following dynamic panel data model:

$$
Y_{it} = \delta_i + \beta_i t + \sum_{k=1}^{K} \alpha_{ik} X_{kt} + \sum_{q=1}^{Q} \gamma_{iq} X_{iqt} + \varepsilon_{it},
$$

$$
\varepsilon_{it} = \varphi_i \varepsilon_{i,t-1} + u_{it},
$$

- $i = 1, ..., N$  denotes the cross-sectional units of the panel (e.g. countries)
- $t = 1, ..., T$  denotes the time period
- $Y_{it}$  denotes the economic or financial dependent variable (e.g. gross domestic product)
- $X_{kt}$  is a set of K exogenous explanatory factors or covariates (e.g. a global volatility index, a market or a commodity index) which are common for all cross-sectional units of the panel, and affect the dependent variables of the panel through separate/different coefficients  $\alpha_{ik}$
- $X_{iqt}$  is a set of Q of cross-sectional specific factors (explanatory variables) which are different for each cross-sectional unit with separate coefficients  $\gamma_{iq}$
- $\delta_i$  are the intercept coefficients of the model, which are different for each cross sectional unit
- $\beta_i$  are the trend coefficients, which are different for each cross sectional unit
- $\varepsilon_{it}$  is a zero-mean autoregressive one, AR(1), process which captures the dynamics of the panel data model. The error term  $u_{it}$  is assumed not to be serially correlated, i.e.  $E(u_{is}u_{it}) = 0$ , for all  $t \neq s$  and for all i, but it is heterogeneous and correlated across i, i.e.  $E(u_{it}u_{jt}) \neq 0$  for all i and  $i$ , and therefore, allows for cross sectional dependence across units

### **Basic linear panel data model**

 The **basic linear panel data model** used in econometric literature can be described through suitable restrictions of the following general model:

 $Y_{it} = \alpha_{it} + \beta_{it} X_{it} + u_{it}$ 

- $i = 1, ..., N$  denotes the cross-sectional units of the panel
- $t = 1, ..., T$  denotes the time period
- $Y_{it}$  denotes the economic or financial dependent variable
- $\bullet$   $X_{it}$  is the explanatory factors or covariates
- $\bullet$   $\alpha_{it}$  is the intercept, which is different for each cross sectional unit, and across time
- $\beta_{it}$  are the explanatory variables coefficients, which are different for each cross sectional unit, and across time
- $u_{it}$  is the error term
- Number of observations: NT
- Number of parameters to be estimated: 2NT
- Can not be estimated !! Thus, **suitable restrictions are imposed** on  $\alpha_{it}$  and  $\beta_{it}$

### **A simple panel data model (model I)**

 The simplest panel data model assumes parameter homogeneity. In the general linear panel model  $Y_{it} = \alpha_{it} + \beta_{it} X_{it} + u_{it}$  impose the following restrictions:  $\alpha_{it} = \alpha$  and  $\beta_{it} = \beta$  for all *i*, *t*. The resulting model can be written in the form:

$$
Y_{it} = a + \beta X_{it} + u_{it},
$$

- $i = 1, ..., N$  denotes the cross-sectional units of the panel
- $t = 1, ..., T$  denotes the time period
- $Y_{it}$  denotes the economic or financial dependent variable
- $\bullet$  X<sub>it</sub> is the explanatory factors or covariate
- $a$  is the intercept of the model, which is common for all cross sectional units and across time
- β is the explanatory variable coefficient, which is common for all cross sectional units across time
- $u_{it}$  is the error term
- This simple panel model does not account for the heterogeneity of the cross sectional units (common  $\alpha$  and  $\beta$  parameter)
- Number of observations: NT
- Number of parameters to be estimated: 2

### **Fixed effects panel data model (model II)**

In the general linear panel data model  $Y_{it} = \alpha_{it} + \beta_{it} X_{it} + u_{it}$  impose the following restrictions:  $\alpha_{it} = \alpha_i$  for all t, and  $\beta_{it} = \beta$  for all i and t. The resulting model is called the fixed effects **model**:

$$
Y_{it} = \alpha_i + \beta X_{it} + u_{it},
$$

- $i = 1, ..., N$  denotes the cross-sectional units of the panel
- $t = 1, ..., T$  denotes the time period
- $Y_{it}$  denotes the economic or financial dependent variable
- $\bullet$  X<sub>it</sub> is the explanatory factors or covariate
- $\bullet$   $\alpha_i$  is the intercept of the model, which is different for each cross sectional unit
- β is the slope, which is common for all cross sectional units
- $u_{it}$  is the error term
- This panel model accounts for the heterogeneity of the cross sectional units (different intercept coefficients  $\alpha_i$ , common  $\,\beta$  parameter)
- Number of observations: NT
- Number of parameters to be estimated: N+1

### **Fixed effects panel data model (model III)**

In the general linear panel data model  $Y_{it} = \alpha_{it} + \beta_{it} X_{it} + u_{it}$  impose the following restrictions:  $\alpha_{it} = \alpha_i$ , and  $\beta_{it} = \beta_i$  for all t. The resulting model is:

 $Y_{it} = \alpha_i + \beta_i X_{it} + u_{it}$ 

- $i = 1, ..., N$  denotes the cross-sectional units of the panel
- $t = 1, ..., T$  denotes the time period
- $Y_{it}$  denotes the economic or financial dependent variable
- $\bullet$   $X_{it}$  is the explanatory factors or covariate
- $\bullet$   $\alpha_i$  is the intercept of the model, which is different for each cross sectional unit
- $\bullet$   $\beta_i$  is the slope, which is different for each cross sectional unit
- $u_{it}$  is the error term
- This panel model accounts for the heterogeneity of the cross sectional units (different intercept coefficients  $\alpha_i$ , and different  $\beta_i$  coefficients)
- Different parameters across units, but constant across time
- Number of observations:
- Number of parameters to be estimated: 2N
- To be able to estimate this model: T>2

### **Least Squares Dummy variables (LSDV)**

Consider the following fixed effects panel data model:

 $Y_{it} = \alpha_i + \beta X_{it} + u_{it}$ 

- Fixed parameters  $\alpha_i$  (non stochastic), different for each sectional unit, common  $\beta$
- We are interested in:
	- Estimate the model parameters
	- Perform hypothesis testing (for the heterogeneity of intercept parameters)
- Re-write the model by using/constructing appropriate dummy variables

$$
Y_{it} = \alpha + \beta X_{it} + \gamma_2 d_{2t} + \gamma_3 d_{3t} + \dots + \gamma_N d_{Nt} + u_{it}
$$

- Number of observations: NT, number of parameters:  $2 + (N 1) = N + 1$
- The model implies that:  $i = 1: Y_{1t} = \alpha + \beta X_{1t} + u_{1t}$  $i = 2$ :  $Y_{2t} = \alpha + \beta X_{2t} + \gamma_2 d_{2t} + u_{2t} \Rightarrow Y_{2t} = (\alpha + \gamma_2) + \beta X_{2t} + u_{2t}$  $i = 3: Y_{3t} = \alpha + \beta X_{3t} + \gamma_3 d_{3t} + u_{3t} \Rightarrow Y_{3t} = (\alpha + \gamma_3) + \beta X_{3t} + u_{3t}$  $\vdots$  $i = N: Y_{Nt} = \alpha + \beta X_{Nt} + \gamma_N d_{Nt} + u_{Nt} \Rightarrow Y_{Nt} = (\alpha + \gamma_N) + \beta X_{Nt} + u_{Nt}$

### **Construct the dummy variables**



Example: number of units  $N = 3$  (i.e. construct  $N-1=2$  dummy variables), time period T, dependent variable  $Y_{it}$ , explanatory variable  $X_{it}$ 

### **Hypothesis testing (t-test)**

For the fixed effects panel data model of the form:

$$
Y_{it} = \alpha + \beta X_{it} + \gamma_2 d_{2t} + \gamma_3 d_{3t} + \dots + \gamma_N d_{Nt} + u_{it}
$$

we can perform two types of hypothesis tests:

**(i) t-test:**

$$
H_0: \gamma_i = 0
$$
  

$$
H_1: \gamma_i \neq 0
$$

by using the following test statistic:

$$
T = \frac{\hat{\gamma}_i}{se(\hat{\gamma}_i)}
$$

The hypothesis testing about parameters  $\gamma_i$  is very important, since if we reject the null hypothesis (H<sub>0</sub>:  $γ<sub>i</sub> = 0$ ), this implies that parameter  $γ<sub>i</sub>$  is statistically significant at level α, and the corresponding intercept parameter for unit *i*, is  $\alpha + \gamma_i$ , and is statistically different than the intercept  $(\alpha)$  of the baseline cross sectional unit 1.

### **Hypothesis testing (F-test)**

#### **(ii) F-test:**

 $H_0: \gamma_2 = \gamma_3 = \cdots = \gamma_N = 0$  $H_1$ : not  $H_0$ The null hypothesis implies the following restricted model:  $H_0$ :  $Y_{it} = \alpha + \beta X_{it} + u_{it}$ 

 while the alternative hypothesis implies the unrestricted model:  $H_1: Y_{it} = \alpha + \beta X_{it} + \gamma_2 d_{2t} + \gamma_3 d_{3t} + \cdots + \gamma_N d_{Nt} + u_{it}$ 

The F-test statistic can be used:

$$
F = \frac{(RSS_R - RSS_{Unr})/(df_R - df_{Unr})}{RSS_{Unr}/df_{Unr}} \sim F_{N-1, NT-N-1}
$$

- RSS<sub>R</sub> and  $RSS_{Unr}$  are the residual sum of squares of the restricted and the unrestricted model, respectively
- $\bullet$   $df_R = NT 2$  and  $df_{IInr} = NT 2 (N 1) = NT N 1$  are the degrees of freedom of the restricted and the unrestricted model, respectively, and  $df_R - df_{Unr} = [NT - 2] [NT - N - 1] = N - 1$

### **Within Estimator**

• Consider the following fixed effects panel data model:

$$
Y_{it} = \alpha_i + \beta X_{it} + u_{it} \quad (1)
$$

Define

$$
\bar{Y}_i = \frac{1}{T} \sum_{t=1}^T Y_{it}, \qquad \bar{X}_i = \frac{1}{T} \sum_{t=1}^T X_{it}, \qquad \bar{u}_i = \frac{1}{T} \sum_{t=1}^T u_{it}
$$

Then

$$
Y_{it} = \alpha_i + \beta X_{it} + u_{it} \Rightarrow
$$
  
\n
$$
\Rightarrow \frac{1}{T} \sum_{t=1}^{T} Y_{it} = \frac{1}{T} \sum_{t=1}^{T} \alpha_i + \frac{1}{T} \beta \sum_{t=1}^{T} X_{it} + \frac{1}{T} \sum_{t=1}^{T} u_{it} \Rightarrow
$$
  
\n
$$
\Rightarrow \overline{Y}_i = \frac{1}{T} T \alpha_i + \beta \overline{X}_i + \overline{u}_i \Rightarrow
$$
  
\n
$$
\Rightarrow \overline{Y}_i = \alpha_i + \beta \overline{X}_i + \overline{u}_i \quad (2)
$$

### **Within Estimator**

By taking the difference (1)-(2) in equations (1) and (2)

$$
Y_{it} = \alpha_i + \beta X_{it} + u_{it} \quad (1)
$$

and

$$
\overline{Y}_i = \alpha_i + \beta \overline{X}_i + \overline{u}_i \ (2)
$$

we obtain:

$$
Y_{it} - \overline{Y}_i = \beta(X_{it} - \overline{X}_i) + (u_{it} - \overline{u}_i)
$$

$$
Y_{it}^* = \beta X_{it}^* + u_{it}^* \quad (3)
$$

- Within Estimator steps:
	- Apply OLS in equation (3) and estimate  $\beta$ , i.e. obtain  $\widehat{\beta}$
	- Using equation (2), estimate  $\alpha_i$ , i.e.  $\widehat{\alpha}_i = \overline{Y}_i \widehat{\beta} \overline{X}_i$
	- Computationally easy
	- Can not conduct hypothesis testing

### **First-difference Estimator**

 Another way of estimating panel data models is by first-differencing the data: lagging the model and subtracting, the time-invariant components are eliminated, and the model

$$
\Delta Y_{i,t} = \beta \Delta X_{i,t} + \Delta u_{i,t}
$$

can be consistently estimated by pooled OLS. This is called the first-difference estimator

- The differences are defined as follows  $\Delta Y_{i,t} = Y_{i,t} Y_{i,t-1}$ ,  $\Delta X_{i,t} = X_{i,t} X_{i,t-1}$ ,  $\Delta u_{i,t} = u_{i,t}$  $u_{i,t-1}$ , for  $t = 2, ..., T$
- Its relative efficiency, and so reasons for choosing it against other consistent alternatives, depends on the properties of the error term. The first-difference estimator is usually preferred if the errors  $u_{it}$  are strongly persistent in time, because then the  $\Delta u_{it}$  will tend to be serially uncorrelated.

## **Panel data models: Random Effects**

### **The Random effects model**

Consider the following panel data model:

$$
Y_{it} = \alpha_i + \beta X_{it} + u_{it} \quad (1)
$$

 $u_{it}$ ~ $iid(0, \sigma^2)$ 

- Suppose that  $a_i \sim D(a, \omega^2)$ , that is:  $E(a_i) = a$  and  $V(a_i) = \omega^2$
- Then, that  $a_i = a + \mu_i$  (2),  $\mu_i \sim iid \ D(0, \omega^2)$ , and  $E(\mu_i) = 0$ , and  $V(\mu_i) = \omega^2$
- In this model, we have two parameters  $(a, \omega^2)$  instead of N parameters  $(a, N 1)$  dummy variables) with respect to the intercept
- The model can be written:

$$
Y_{it} = \alpha_i + \beta X_{it} + u_{it} \Rightarrow
$$
  
\n
$$
\Rightarrow Y_{it} = a + \mu_i + \beta X_{it} + u_{it} \Rightarrow
$$
  
\n
$$
\Rightarrow Y_{it} = a + \beta X_{it} + (\mu_i + u_{it}) \Rightarrow
$$
  
\n
$$
\Rightarrow Y_{it} = a + \beta X_{it} + v_{it}, \text{ where } v_{it} = \mu_i + u_{it}
$$

## **Panel data models: Random Effects**

### **The Random effects model**

For the process

 $v_{it} = \mu_i + u_{it}$ 

we observe that:  $v_{i1} = \mu_i + u_{i1}$ ,  $v_{i2} = \mu_i + u_{i2}$ , …,  $v_{iT} = \mu_i + u_{iT}$ . Therefore,  $v_{it}$  contains a common component,  $\mu_i$ , across time, and thus  $v_{it}$  is auto-correlated.

• Its mean is: 
$$
E(v_{it}) = E(\mu_i + u_{it}) = E(\mu_i) + E(u_{it}) = 0
$$

The variance is:

$$
V(v_{it}) = V(\mu_i + u_{it}) = V(\mu_i) + V(u_{it}) = \omega^2 + \sigma^2
$$

• The covariance of  $v_{it}$  with  $v_{is}$ ,  $t \neq s$ , is:

$$
Cov(v_{it}, v_{is}) = E(v_{it}v_{is}) = E[(\mu_i + u_{it})(\mu_i + u_{is})] =
$$
  

$$
= E[\mu_i^2 + \mu_i u_{is} + u_{it}\mu_i + u_{it}u_{is}] =
$$
  

$$
= E[\mu_i^2] = V(\mu_i) = \omega^2 \neq 0, \text{ there is auto-correlation at } v_{it}
$$

### **Panel data models: Random Effects**

### **The Random effects model**

• Therefore, the covariance matrix of  $v_i = (v_{i1}, v_{i2}, ..., v_{iT})'$  is

$$
\Omega = \begin{bmatrix} \omega^2 + \sigma^2 & \omega^2 & \cdots & \omega^2 \\ \omega^2 & \omega^2 + \sigma^2 & \cdots & \omega^2 \\ \vdots & \vdots & \ddots & \vdots \\ \omega^2 & \omega^2 & \cdots & \omega^2 + \sigma^2 \end{bmatrix}
$$

 However, there is not heteroskedasticity. The covariances between the elements of  $v_i = (v_{i1}, v_{i2}, ..., v_{iT})'$  and  $v_j = (v_{j1}, v_{j2}, ..., v_{jT})'$  are given by:

$$
Cov(v_{it}, v_{jt}) = E(v_{it}v_{jt}) = E[(\mu_i + u_{it})(\mu_j + u_{jt})] =
$$

$$
= E[\mu_i \mu_j + \mu_i u_{jt} + u_{it} \mu_j + u_{it} u_{jt}] = 0
$$

Thus,

$$
Cov\begin{pmatrix} v_1 \\ v_2 \\ \vdots \\ v_N \end{pmatrix} = \begin{bmatrix} \Omega & 0 & \cdots & 0 \\ 0 & \Omega & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \Omega \end{bmatrix}
$$

### **Fixed or Random Effects Model (Hausman test)**

- **Decide between fixed or random effects model: we can run the Hausman test**
- Null hypothesis: the preferred model is random effects vs. the alternative the fixed effects (see Green, 2008, chapter 9).
- Steps:
	- Run a fixed effects model and save the estimates
	- Run a random model and save the estimates
	- Perform Hausman test. If the p-value is significant (for example a=0.05 > p-value) then use fixed effects, if not use random effects
- R command:
	- fixed <-  $plm(y \sim sf1, data=data, index=c("country"), model="within")$
	- random  $\leq$  plm(y  $\approx$  sf1, data=data, index=c("country"), model="random")
	- phtest(fixed, random)

### **Tests - Diagnostics**

- **Breusch-Pagan Lagrange Multiplier for random effects**
- The LM test helps you decide between a random effects regression and a simple OLS regression
- The null hypothesis in the LM test: no panel effect (i.e. OLS better). That is, no significant difference across units (i.e. no panel effect)
- R command:
	- pool <-  $pIm(y \sim sf1, data = data, model="pooling")$
	- plmtest(pool, type=c("bp"))

### **Tests - Diagnostics**

- **Breusch-Pagan Lagrange Multiplier test for cross-sectional dependence in panels**
- Cross-sectional dependence is a problem in panel data (especially with long time series)
- The null hypothesis in the Breusch-Pagan/LM Cross-sectional dependence tests is that residuals across units are not correlated
- Breusch-Pagan/LM (cross-sectional dependence) tests are used to test whether the residuals are correlated across units
- R commands:
	- Breusch-Pagan Lagrange Multiplier test for cross-sectional dependence in panels
	- fixed <-plm(y  $\sim$  sf1, data=data, index=c("country"), model="within")
	- pcdtest(fixed, test =  $c("Im")$ )
	- Pesaran test for cross-sectional dependence in panels
	- pcdtest(fixed, test =  $c("cd")$ )

### **Testing for serial correlation**

- **Breusch-Godfrey/Wooldridge test for serial correlation in panel models**
- Serial correlation tests apply to panel data. The null is that there is not serial correlation
- R commands: (required package: lmtest)
	- fixed <-plm(y  $\sim$  sf1, data=data, index=c("country"), model="within")
	- pbgtest(fixed)

### **Testing for heteroskedasticity**

- **Breusch-Pagan test for heteroskedasticity**
- The null hypothesis for the Breusch-Pagan test is homoskedasticity
- If hetersokedaticity is detected we can use robust covariance matrix to account for it, or model the conditional variances (using, for example, ARCH/GARCH type models)
- R commands: (required package: lmtest)
	- o bptest(y  $\sim$  sf1 + factor(country), data = data, studentize=F)

### **Controlling for heteroskedasticity**

- Robust covariance matrix estimation (Sandwich estimator)
- The 'vcovHC' function estimates three heteroskedasticity-consistent covariance estimators:
	- "white1": for general heteroskedasticity but no serial correlation (recommended for random effects)
	- "white2": is "white1" restricted to a common variance within groups (recommended for random effects)
	- "arellano": both heteroskedasticity and serial correlation (recommended for fixed effects)
- The following options can be applied:
	- HC0 : heteroskedasticity consistent (default)
	- HC1,HC2, HC3 : Recommended for small samples. HC3 gives less weight to influential observations
	- HC4 : small samples with influential observations
	- HAC : heteroskedasticity and autocorrelation consistent (type ?vcovHAC for more details)
- R commands: (required package: lmtest)
	- fixed  $\le$ -plm(y  $\approx$  sf1, data=data, index=c("country"), model="within")
	- coeftest(fixed) # Original coefficients
	- coeftest(fixed, vcovHC) # Heteroskedasticityconsistent coefficients
	- coeftest(fixed, vcovHC(fixed, method = "arellano")) # Heteroskedasticity consistent coefficients (Arellano)
	- random <-plm( $y \sim$  sf1, data=data, index=c("country"), model="random")
	- coeftest(random) # Original coefficients
	- coeftest(random, vcovHC) # Heteroskedasticity consistent coefficients

# **Panel data models: Application to R**

- Several panel data models will be implemented in R
- See corresponding R-file

# **Thank you**

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