

## An example of a linear model with instrumental variables

Consider the process  $(y_t)_{t \in \mathbb{Z}}$  defined by  $y_t = b_0 x_t + \epsilon_t$ ,  $b_0 \in \mathbb{R}$ , where  $(\epsilon_t)_{t \in \mathbb{Z}}$  is a strictly stationary and ergodic white noise process ( $WN(\sigma^2)$ ) and  $(x_t)_{t \in \mathbb{Z}}$  is a linear process based on the white noise process  $(\epsilon_t)_{t \in \mathbb{Z}}$  thus for a real sequence  $(\alpha_j)_{j \in \mathbb{N}}$ ,

$$\sum_{j=0}^{\infty} |\alpha_j| < +\infty, \alpha_j, \alpha_{j+1} \neq 0 \text{ for some } j \in \mathbb{N}, x_t = \sum_{j=0}^{\infty} \alpha_j \epsilon_{t-j}.$$

Given a sample  $(y_t, x_t)_{t=1, \dots, T}$ , suppose that our interest centers on estimating the unknown  $b_0$ . Propose an estimator for  $b_0$  and derive its asymptotic properties.

Let's consider the OLS estimator for  $b_0$ :

$$b_T = \frac{\sum_{t=1}^T x_t y_t}{\sum_{t=1}^T x_t^2} = \frac{\sum_{t=1}^T x_t (b_0 x_t + \epsilon_t)}{\sum_{t=1}^T x_t^2} = b_0 + \frac{\frac{1}{T} \sum_{t=1}^T x_t \epsilon_t}{\frac{1}{T} \sum_{t=1}^T x_t^2}.$$

Given the results in the paragraph Transformations and Heredity and the ones about Causal Linear Processes, since  $(\epsilon_t)_{t \in \mathbb{Z}}$  is strictly stationary and ergodic, we have that  $(x_t)_{t \in \mathbb{Z}}$ ,  $(x_t^2)_{t \in \mathbb{Z}}$ ,  $(x_t \epsilon_t)_{t \in \mathbb{Z}}$  are strictly stationary and ergodic processes.

Also,

$$\begin{aligned} E(x_t \epsilon_t) &= E\left[\left(\sum_{j=0}^{\infty} \alpha_j \epsilon_{t-j}\right) \cdot \epsilon_t\right] = E\left(\sum_{j=0}^{\infty} \alpha_j \epsilon_{t-j} \epsilon_t\right) = \sum_{j=0}^{\infty} \alpha_j E(\epsilon_{t-j} \epsilon_t) \\ &= \alpha_0 E(\epsilon_t^2) + \sum_{j=1}^{\infty} \alpha_j E(\epsilon_{t-j} \epsilon_t) = \alpha_0 \sigma^2 < +\infty. \end{aligned} \quad \text{since } \epsilon \text{ is white noise}$$

Obviously, we have an endogeneity issue here, as the error term  $\epsilon_t$  is not orthogonal to the regressor  $x_t$ .

Moreover,

$$E(x_t^2) = E\left[\left(\sum_{j=0}^{\infty} \alpha_j \epsilon_{t-j}\right)^2\right] = E\left[\sum_{j=0}^{\infty} \sum_{j^*=0}^{\infty} \alpha_j \alpha_{j^*} \epsilon_{t-j} \epsilon_{t-j^*}\right] = \sum_{j, j^*=0}^{\infty} \alpha_j \alpha_{j^*} E(\epsilon_{t-j} \epsilon_{t-j^*})$$

$$\text{since } \epsilon \text{ is white noise } E(\epsilon_{t-j} \epsilon_{t-j^*}) = \begin{cases} 0, & j \neq j^* \\ \sigma^2, & j = j^* \end{cases} \text{ hence,}$$

$$E(x_t^2) = \sum_{j, j^*=0}^{\infty} \alpha_j^2 E(\epsilon_{t-j}^2) = \sigma^2 \sum_{j=0}^{\infty} \alpha_j^2 < +\infty. \quad \text{since } \sum_{j=0}^{\infty} |\alpha_j| < +\infty.$$

Recall, Birkhoff's Law of Large Numbers:

Suppose that  $(x_t)_{t \in \mathbb{Z}}$  is stationary ergodic, and that  $E(|x_0|) < +\infty$  exists as a real number. Then,  $\lim_{T \rightarrow \infty} \frac{1}{T} \sum_{t=0}^T x_t \xrightarrow{a.s.} E(x_0)$  P a.s.

Thus, by Birkhoff's LLN:

$$\cdot \frac{1}{T} \sum_{t=0}^T x_t \epsilon_t \xrightarrow{a.s.} E(x_0 \epsilon_0) = \alpha_0 \sigma^2 \quad \text{P a.s.}$$

$$\cdot \frac{1}{T} \sum_{t=0}^T x_t^2 \xrightarrow{a.s.} E(x_0^2) = \sigma^2 \sum_{j=0}^{\infty} \alpha_j^2 > 0 \quad \text{P a.s.}$$

and thereby due to the Continuous Mapping Theorem:

$$b_T = b_0 + \frac{\frac{1}{T} \sum_{t=0}^T x_t \epsilon_t}{\frac{1}{T} \sum_{t=0}^T x_t^2} \xrightarrow{a.s.} b_0 + \underbrace{\frac{\alpha_0 \sigma^2}{\sigma^2 \sum_{j=0}^{\infty} \alpha_j^2}}_{\neq 0} \quad \text{P a.s.}$$

Therefore, the OLS estimator for  $b_0$  is inconsistent.

To overcome the endogeneity problem we use the so-called instrumental variables estimator (IV). This approach assumes the existence of a variable  $z$  which we call instrument, that is correlated with the endogenous variable  $\text{Cov}(z_t, x_t) \neq 0$  but exogenous to the error term i.e.  $\text{Cov}(z_t, \epsilon_t) = 0$ .

Given the structure of  $x_t$ , a valid instrument in our case would be  $x_{t-1}$  defined by  $x_{t-1} = \sum_{j=0}^{\infty} \alpha_j \epsilon_{t-1-j}$ , since:

$$\cdot \text{Cov}(x_t, x_{t-1}) = E(x_t x_{t-1}) = \sigma^2 \sum_{j=0}^{\infty} \alpha_j \alpha_{j+1} < +\infty$$

given that we have already proved that for a linear process  $(x_t)_{t \in \mathbb{Z}}$ ,  $\text{Cov}(x_t, x_{t-k}) = E(\epsilon_0^2) \sum_{j=0}^{\infty} \alpha_j \alpha_{j+k}$  and  $\sum_{j=0}^{\infty} |\alpha_j \alpha_{j+k}| < +\infty$ .

$$\cdot \text{Cov}(x_{t-1}, \epsilon_t) = E(x_{t-1} \epsilon_t) = E\left[\left(\sum_{j=0}^{\infty} \alpha_j \epsilon_{t-1-j}\right) \epsilon_t\right] = \sum_{j=0}^{\infty} \alpha_j E(\epsilon_{t-1-j} \epsilon_t) = 0$$

since  $(\epsilon_t)_{t \in \mathbb{Z}}$  is a WN.

Hence,  $x_{t-1}$  is a valid instrument which enables the definition of a moment condition which along with an identification condition and the analogy principle, give us enough structure so as a criterion function can be constructed, upon the optimization of which, we can design the IV estimation (see Tutorial - Linear model with instrumental variables).

Remember that the IV estimator is denoted by :

$$\begin{aligned} b_{IV} &= (Z'X)^{-1} Z'Y = (Z'X)^{-1} Z'(Xb_0 + \epsilon) = b_0 + (Z'X)^{-1} Z'\epsilon \\ &= b_0 + \left( \frac{1}{T} \sum_{t=1}^T z_t x_t \right)^{-1} \frac{1}{T} \sum_{t=1}^T z_t \epsilon_t \end{aligned}$$

where  $z_t$  in our case is  $x_{t-1}$ , therefore given that we also know  $x_0$ ,

$$b_{IV} = b_0 + \left( \frac{1}{T} \sum_{t=1}^T x_{t-1} x_t \right)^{-1} \frac{1}{T} \sum_{t=1}^T x_{t-1} \epsilon_t$$

Given the results in the paragraph Transformations and Heridicity, since  $(\epsilon_t)_{t \in \mathbb{Z}}$  is stationary ergodic, we have that  $(x_{t-1} x_t)_{t \in \mathbb{Z}}$  and  $(x_{t-1} \epsilon_t)_{t \in \mathbb{Z}}$  are stationary ergodic processes. Also,  $E(x_t x_{t-1}) < +\infty$  and  $E(x_{t-1} \epsilon_t) < +\infty$ . Hence, by Birkhoff's LLN :

$$\bullet \frac{1}{T} \sum_{t=1}^T x_{t-1} x_t \xrightarrow{\text{P a.s.}} E(x_{-1} x_0) = \sigma^2 \sum_{j=0}^{\infty} a_j a_{j+1}$$

$$\bullet \frac{1}{T} \sum_{t=1}^T x_{t-1} \epsilon_t \xrightarrow{\text{P a.s.}} E(x_{-1} \epsilon_0) = 0$$

Since  $\sum_{j=0}^{\infty} a_j a_{j+1} \neq 0$  since  $a_j, a_{j+1} \neq 0$  for some  $j \in \mathbb{N}$ , due to the CMT :

$$b_{IV} = b_0 + \frac{\frac{1}{T} \sum_{t=1}^T x_{t-1} \epsilon_t}{\frac{1}{T} \sum_{t=1}^T x_{t-1} x_t} \xrightarrow{\text{P a.s.}} b_0 + \frac{0}{\sigma^2 \sum_{j=0}^{\infty} a_j a_{j+1}} = b_0$$

Therefore, the IV estimator is strongly consistent

Consider now that we want to find the asymptotic distribution of the proposed IV estimator.

$$\sqrt{T}(\hat{\beta}_{IV} - \beta_0) = \left( \frac{1}{T} \sum_{t=1}^T x_{t-1} x_t \right)^{-1} \frac{1}{\sqrt{T}} \sum_{t=1}^T x_{t-1} \epsilon_t$$

We have already specified the limit of  $\frac{1}{T} \sum_{t=1}^T x_{t-1} x_t$ , therefore the limit distribution is going to be determined by  $\frac{1}{\sqrt{T}} \sum_{t=1}^T x_{t-1} \epsilon_t$ .

We are examining the properties of  $(x_{t-1} \epsilon_t)_{t \in \mathbb{Z}}$  process and we have already said that is stationary ergodic, thus we need to investigate whether it is a square martingale difference sequence (s.m.d.) or not in order to use the appropriate Central Limit Theorem.

Remember that:

The process  $(x_t)_{t \in \mathbb{Z}}$  is called a square martingale difference sequence wrt. the filtration  $(f_t)_{t \in \mathbb{Z}}$ , iff

1.  $(x_t)_{t \in \mathbb{Z}}$  is adapted to  $(f_t)_{t \in \mathbb{Z}}$
2.  $E(x_t^2) < +\infty \quad \forall t \in \mathbb{Z}$
3.  $E(x_t | f_{t-1}) = 0 \quad \forall t \in \mathbb{Z}$

Strict stationarity, ergodicity and the s.m.d. property imply the following CLT:

Suppose that  $(x_t)_{t \in \mathbb{Z}}$  is strictly stationary and ergodic, and that for some  $(f_t)_{t \in \mathbb{Z}}$ ,  $((x_t)_{t \in \mathbb{Z}}, (f_t)_{t \in \mathbb{Z}})$  is s.m.d.

Then,  $\frac{1}{\sqrt{T}} \sum_{t=1}^T x_t \xrightarrow{d} \mathcal{N}(0, E(x_0^2))$ , as  $T \rightarrow \infty$ .

In our case consider the filtration  $f_t := \sigma(\epsilon_{t-i}, i \geq 0)$  and further assume that  $(\epsilon_t)_{t \in \mathbb{Z}}$  is s.m.d. wrt some filtration

- $(x_{t-1} \epsilon_t)_{t \in \mathbb{Z}}$  is adapted to  $f_t$  by construction
- $E(x_{t-1} \epsilon_t / f_{t-1}) = x_{t-1} E(\epsilon_t / f_{t-1}) = x_{t-1} \cdot 0 = 0 \quad \forall t \in \mathbb{Z}$

since  $(\epsilon_t)_{t \in \mathbb{Z}}$  is s.m.d.

$$E[(x_{t-1} \epsilon_t)^2] = E(x_{t-1}^2 \epsilon_t^2) = E\left[\left(\sum_{j=0}^{\infty} a_j \epsilon_{t-1-j}\right)^2 \epsilon_t^2\right]$$

$$= E\left[\left(\sum_{j=0}^{\infty} a_j^2 \epsilon_{t-1-j}^2 + \sum_{\substack{i,j=0 \\ i \neq j}}^{\infty} a_i a_j \epsilon_{t-1-i} \epsilon_{t-1-j}\right) \epsilon_t^2\right]$$

$$\begin{aligned}
&= \sum_{j=0}^{\infty} \alpha_j^2 E(\epsilon_t^2 \epsilon_{t-1-j}^2) + \sum_{\substack{i,j=0 \\ i \neq j}}^{\infty} \alpha_i \alpha_j E(\epsilon_t^2 \epsilon_{t-1-j} \epsilon_{t-1-i}) \\
&\leq \sum_{j=0}^{\infty} \alpha_j^2 E(\epsilon_t^2 \epsilon_{t-1-j}^2) + \left| \sum_{\substack{i,j=0 \\ i \neq j}}^{\infty} \alpha_i \alpha_j E(\epsilon_t^2 \epsilon_{t-1-j} \epsilon_{t-1-i}) \right|
\end{aligned}$$

However, no already stated assumptions specify that the above sums converge.

Therefore, we need to assume that  $\sum_{j=0}^{\infty} \alpha_j^2 E(\epsilon_t^2 \epsilon_{t-1-j}^2) < +\infty$  and

$$\left| \sum_{\substack{i,j=0 \\ i \neq j}}^{\infty} \alpha_i \alpha_j E(\epsilon_t^2 \epsilon_{t-1-j} \epsilon_{t-1-i}) \right| < +\infty.$$

Hence, the process  $(x_{t-1}, \epsilon_t)_{t \in \mathbb{Z}}$  is s.m.d and we can use the relevant CLT.

$$\frac{1}{\sqrt{T}} \sum_{t=1}^T x_{t-1} \epsilon_t \xrightarrow{d} \mathcal{N}(0, E(x_{t-1}^2 \epsilon_0^2)) \text{ as } T \rightarrow \infty.$$

and due to the CMT

$$\sqrt{T} (b_{IV} - b_0) = \left( \frac{1}{T} \sum_{t=1}^T x_{t-1} x_t \right)^{-1} \frac{1}{\sqrt{T}} \sum_{t=1}^T x_{t-1} \epsilon_t \xrightarrow{d} [E(x_{t-1} x_0)]^{-1} \mathcal{N}\left(0, \frac{E(x_{t-1}^2 \epsilon_0^2)}{[E(x_{t-1} x_0)]^2}\right)$$