

Lecture 3: Identification in Linear Models

Econometrics — *continuation from Lecture 2*

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▷ Amber boxes = Handwritten Notes (professor's words)

◊ Teal boxes = Student's Notes

Recall from Lecture 2

In Lecture 2 we set up the semi-parametric linear model $Y_n = X_n b_0 + \epsilon_n$ with three assumptions (exogeneity, spherical errors, full rank) and proposed the expected conditional squared-error loss

$$U_n^*(b) := \mathbb{E} \left[\frac{1}{n} (Y_n - X_n b)' (Y_n - X_n b) \mid X_n \right]$$

as our objective function. **This lecture proves that U_n^* is uniquely minimised at the true parameter b_0** —i.e. the identification result.

1 The Core Intuition

▷ Handwritten Notes (what the professor said)

We want to uncover the true relationship, b_0 , that generated our data. If we guess a parameter b , how wrong is our guess on average? The mathematical proof below demonstrates that our “average squared mistake” forms a perfect, multi-dimensional *bowl*.

Because of our statistical assumptions, the absolute bottom of this bowl occurs exactly when our guess matches reality ($b = b_0$). Keeping this “bowl” shape in mind helps anchor the algebra: every step is simply moving toward proving that the loss function is a parabola pointing upwards.

◇ Student's Notes

Think of the scalar analogy first. If $p = 1$ the expected loss becomes

$$U_n^*(b) = \underbrace{\sum_{i=1}^n x_i^2}_{>0} (b_0 - b)^2 + 1,$$

a literal upward-opening parabola in b with vertex at b_0 . The full proof generalises this to p dimensions: the parabola becomes an *elliptic paraboloid* because $\frac{X'X}{n}$ is positive definite.

Road-map of the proof:

1. Substitute the DGP into the residual.
2. Expand the quadratic and take the conditional expectation.
3. Show the cross-term dies (exogeneity) and the constant term equals 1 (spherical errors).
4. Conclude uniqueness via positive definiteness of $X'X/n$.

2 Model Setup (Recap from Lecture 2)

▷ Handwritten Notes (what the professor said)

Our data generating process is defined as:

$$Y_n = X_n b_0 + \epsilon_n$$

The critical assumptions are:

- **Strict Exogeneity:** $\mathbb{E}(\epsilon_n | X_n) = 0_{n \times 1}$ (*The errors are pure noise.*)
- **Spherical Errors:** $\text{Var}(\epsilon_n | X_n) = I_{n \times n}$ (*The noise is uniform and uncorrelated.*)
- **Full Rank (Identification):** $\text{rank}(X_n) = p$ (*Enough unique info to pinpoint the parameters.*)

◇ Student's Notes

Dimension check. Y_n is $n \times 1$, X_n is $n \times p$, b_0 is $p \times 1$, ϵ_n is $n \times 1$.

- *Strict exogeneity* is stronger than just $\text{Cov}(X, \epsilon) = 0$; it says **no column** of X_n carries information about **any** ϵ_i .
- *Spherical errors* bundles two things: homoskedasticity ($\text{Var}(\epsilon_i | X) = 1$ for every i)

and no serial correlation ($\text{Cov}(\epsilon_i, \epsilon_j|X) = 0$ for $i \neq j$). When these fail, we move to GLS.

- *Full rank* requires $n \geq p$ (at least as many observations as parameters) **and** no exact multicollinearity. This is what makes $X'X$ invertible.

! Watch Out

The spherical-errors assumption here sets the variance to $\sigma^2 I$ with $\sigma^2 = 1$ for notational simplicity. In the general case $\text{Var}(\epsilon|X) = \sigma^2 I$, every formula carries a σ^2 factor, but the identification argument is identical.

3 The Objective Function (The “Mistake” Function)

▷ Handwritten Notes (what the professor said)

We define our expected loss function as:

$$U_n^*(b) := \mathbb{E} \left[\frac{1}{n} (Y_n - X_n b)' (Y_n - X_n b) \mid X_n \right]$$

First, substitute the true model into our guess to see the error gap:

$$Y_n - X_n b = X_n b_0 + \epsilon_n - X_n b = X_n (b_0 - b) + \epsilon_n$$

Next, expand the quadratic form algebraically:

$$\begin{aligned} & (X_n (b_0 - b) + \epsilon_n)' (X_n (b_0 - b) + \epsilon_n) \\ &= \underbrace{(b_0 - b)' X_n' X_n (b_0 - b)}_{\text{Term 1}} + \underbrace{2(b_0 - b)' X_n' \epsilon_n}_{\text{Term 2}} + \underbrace{\epsilon_n' \epsilon_n}_{\text{Term 3}} \end{aligned}$$

◇ Student’s Notes

This is just the identity $(a + b)'(a + b) = a'a + 2a'b + b'b$ with

$$a = X_n (b_0 - b), \quad b = \epsilon_n.$$

Each of the three terms will be handled separately under the conditional expectation in the next section.

4 Applying the Expectation (Filtering the Noise)

▷ Handwritten Notes (what the professor said)

Now, we apply the conditional expectation operator $\frac{1}{n} \mathbb{E}(\cdot | X_n)$ to each of the three terms. This is where we mathematically “average out” the noise to reveal the underlying structure.

4.1 Term 1: The Deterministic Signal

▷ Handwritten Notes (what the professor said)

Since b , b_0 , and X_n are treated as constants under the condition, they pass through unchanged:

$$\frac{1}{n} \mathbb{E}[(b_0 - b)' X_n' X_n (b_0 - b) | X_n] = (b_0 - b)' \left(\frac{X_n' X_n}{n} \right) (b_0 - b)$$

◇ Student's Notes

Conditioning on X_n means we treat the entire design matrix as *known*. The only random object inside the expectation is ϵ_n , and this term contains no ϵ_n at all—so the expectation is the identity operation here.

The matrix $\frac{X_n' X_n}{n}$ is the **sample second-moment matrix** of the regressors. It will play the role of the curvature of our bowl.

4.2 Term 2: The Cross-Product (Zeroing Out)

▷ Handwritten Notes (what the professor said)

Because the errors are pure noise ($\mathbb{E}(\epsilon_n | X_n) = 0$), this term vanishes. Recognizing that the expected value of the noise is zero immediately simplifies the expression:

$$\frac{1}{n} \mathbb{E}[2(b_0 - b)' X_n' \epsilon_n | X_n] = \frac{2}{n} (b_0 - b)' X_n' \underbrace{\mathbb{E}(\epsilon_n | X_n)}_{=0} = 0$$

◇ Student's Notes

This is the **crucial** step powered by strict exogeneity. If exogeneity failed (e.g. an omitted variable that correlates with both X and ϵ), this cross-term would *not* vanish and the minimum of U_n^* would *not* be at b_0 —we would have **bias**.

Linearity of expectation lets us pull the non-random $(b_0 - b)' X_n'$ out, leaving only $\mathbb{E}(\epsilon_n | X_n)$ inside.

4.3 Term 3: The Trace Trick (The Error Variance)

▷ Handwritten Notes (what the professor said)

For a scalar value like $\epsilon'_n \epsilon_n$, the value equals its trace. Because the trace is a linear operator, we swap the trace and expectation to isolate the variance matrix:

$$\frac{1}{n} \mathbb{E}[\epsilon'_n \epsilon_n | X_n] = \frac{1}{n} \mathbb{E}[\text{tr}(\epsilon_n \epsilon'_n) | X_n] = \frac{1}{n} \text{tr}(\mathbb{E}[\epsilon_n \epsilon'_n | X_n])$$

Since $\text{Var}(\epsilon_n | X_n) = I_{n \times n}$, the trace of an $n \times n$ identity matrix is simply n :

$$\frac{1}{n} \text{tr}(I_{n \times n}) = \frac{n}{n} = 1$$

◇ Student's Notes

Why the trace trick? $\epsilon'_n \epsilon_n$ is a 1×1 scalar, while $\epsilon_n \epsilon'_n$ is an $n \times n$ matrix. We want the matrix form because we know $\mathbb{E}[\epsilon_n \epsilon'_n | X_n] = \text{Var}(\epsilon_n | X_n)$ (since $\mathbb{E}[\epsilon_n | X_n] = 0$). The cyclic property of trace bridges the two:

$$\epsilon'_n \epsilon_n = \text{tr}(\epsilon'_n \epsilon_n) = \text{tr}(\epsilon_n \epsilon'_n).$$

The first equality holds because any scalar equals its own trace. The second uses $\text{tr}(AB) = \text{tr}(BA)$.

General case: if $\text{Var}(\epsilon | X) = \sigma^2 I$, this term becomes σ^2 instead of 1. It shifts the bowl up but does not change the location of the minimum.

5 The Unique Minimum

▷ Handwritten Notes (what the professor said)

Reassembling the filtered function gives us:

$$U_n^*(b) = (b_0 - b)' \left(\frac{X'_n X_n}{n} \right) (b_0 - b) + 1$$

Because X_n has full rank, the matrix $A = \frac{X'_n X_n}{n}$ is **positive definite**. This mathematically guarantees our “bowl” shape.

For any positive definite matrix A , the quadratic form $y' A y \geq 0$, and it *strictly* equals 0 if and only if $y = 0_{p \times 1}$.

Let $y = (b_0 - b)$. The minimum possible value of this entire function is uniquely achieved when:

$$b_0 - b = 0 \implies b = b_0$$

Therefore, minimizing this objective function cleanly and uniquely identifies the true parameter b_0 :

$$\arg \min_{b \in \Theta} U_n^*(b) = b_0$$

Key Result

$$\arg \min_{b \in \Theta} U_n^*(b) = b_0$$

The expected conditional loss is **uniquely minimised** at the true parameter vector. This is the **identification** result promised at the end of Lecture 2: b_0 is the *only* value in the parameter space that minimises the population objective, so it is in principle recoverable from data.

◇ Student's Notes

Unpacking positive definiteness: A symmetric matrix A is positive definite iff all its eigenvalues $\lambda_1, \dots, \lambda_p > 0$. Since $A = \frac{1}{n}X'X$ and $\text{rank}(X) = p$, the columns of X are linearly independent, which forces every eigenvalue of $X'X$ (and hence A) to be strictly positive.

Geometrically, positive definiteness means the level sets $\{b : U_n^*(b) = c\}$ are **ellipsoids** centred at b_0 ; there are no flat directions, so the minimum is unique.

Why this matters going forward:

1. *Identification* is a **population** concept—it says the truth *can* in principle be found; it does not say how fast or accurately a finite sample does so.
2. The *estimation* step replaces U_n^* with the sample analogue $\hat{U}_n(b) = \frac{1}{n}(Y - Xb)'(Y - Xb)$ and minimises it, giving the OLS estimator $\hat{b} = (X'X)^{-1}X'Y$.
3. *Consistency* then asks whether $\hat{b} \rightarrow b_0$ as $n \rightarrow \infty$ —this is where the law of large numbers (for $X'X/n$ and $X'\epsilon/n$) enters.

Quick-Reference Summary

◇ Student's Notes

Term	Expression	Role
Signal	$(b_0 - b)' \frac{X'X}{n} (b_0 - b)$	Bowl curvature; = 0 iff $b = b_0$
Cross	$\frac{2}{n}(b_0 - b)' X'\epsilon$	Killed by exogeneity
Noise	$\frac{1}{n}\epsilon'\epsilon$	$\rightarrow 1$ (or σ^2); constant shift

Lectures 2–3 narrative arc:

Lecture	What was accomplished
Lect. 2	Defined parametric vs. semi-parametric models; chose the objective function
Lect. 3	Proved the objective has a unique minimum at b_0 (identification)
Next	Replace population objective with sample analogue \Rightarrow OLS estimator