Athens University of Economics and Business *Department of Economics*

Postgraduate Program - Master's in Economic Theory *Course: Econometrics II* Prof: Stelios Arvanitis TA: Alecos Papadopoulos

Semester: Spring 2017-2018

March 22, 2018

An example of a "Long-memory" process

Definition: A weakly stationary stochastic process $(x_t)_{t \in \mathbb{Z}}$ with autocovariance function

 γ_k is said to have "long-memory" if $\sum_{k=0}^{\infty} | \gamma_k | = \infty$.

An example.

Consider $(u_t)_{t \in \mathbb{Z}}$ a strictly stationary White Noise process $WN(\sigma_u^2)$, and a real sequence

 $\left(\theta_{_j}\right)_{_{j\in\mathbb{N}}}$. $\ \forall\, t\!\in\!\mathbb{Z} \,$ we initially define

$$
y_t = \sum_{j=0}^{\infty} \theta_j u_{t-j}, \qquad \theta_j = \frac{\Gamma(j+d)}{\Gamma(j+1)\Gamma(d)}, \quad d > 0,
$$

where $\Gamma(\cdot)$ is the Gamma function. The parameter d is called the "memory parameter". We will see that in order for y_t to be well defined, we will have to strengthen the

restrictions on the value of d to $0 < d < 1/2$.

The Gamma function satisfies the identity

$$
\Gamma(j+1) = j\Gamma(j) \Rightarrow \theta_j = \frac{\Gamma(j+d)}{j\Gamma(j)\Gamma(d)}
$$

The following exact relation holds between the Gamma function and the Beta function

$$
B(j,d) = \frac{\Gamma(j)\Gamma(d)}{\Gamma(j+d)} \Rightarrow \theta_j = \frac{1}{jB(j,d)}
$$

Using the Stirling approximation for "large" j and fixed/small d as is our case, we have

$$
B(j,d) \approx \Gamma(d) j^{-d} \Rightarrow \theta_j \approx \frac{1}{j \Gamma(d) j^{-d}} \Rightarrow \theta_j \approx \frac{1}{\Gamma(d) j^{1-d}}
$$
 while note that $\theta_0 = 1$ exactly.

Note also that $\Gamma(d)$ is bounded and does not depend on the index.

Convergence.

From the theory of infinite series we know that

$$
\sum_{j=1}^{\infty} \frac{1}{j^{1-d}}
$$
 diverges since $1-d \le 1$, $\forall d > 0$

So
$$
\sum_{j=0}^{\infty} |\theta_j| = \sum_{j=0}^{\infty} \theta_j \longrightarrow 1 + \frac{1}{\Gamma(d)} \sum_{j=1}^{\infty} \frac{1}{j^{1-d}} \longrightarrow \infty
$$

And **the coefficient series is not absolutely summable.**

But consider

$$
\sum_{j=0}^{\infty} \theta_j^2 \longrightarrow 1 + \frac{1}{\left(\Gamma(d)\right)^2} \sum_{j=1}^{\infty} \frac{1}{j^{2-2d}}
$$

Whether this converges depends on the actual value of *d* **.**

If
$$
0 < d < 1/2
$$
 we obtain $2 - 2d > 2 - 2\frac{1}{2} = 1$, in which case $\sum_{j=0}^{\infty} \theta_j^2 < +\infty$.

Namely, for $0 < d < 1/2$ we have square-summability of the coefficient series. Square-summability guarantees that 0 $t - \sum_{j}^{\mathbf{U}} t_{j} \mathbf{u}_{t-j}$ *j* $y_i = \sum_i \theta_i u_i$ ∞ ⁻ $=\sum_{j=0}\theta_ju_{i-j}$ "converges in mean-square" to some random variable (see Hamilton's book Appendix 3.A, pp 69-70).

So **we eventually impose** $0 < d < 1/2$, and we have square-summability but not absolute summability. Square-summability (and hence convergence), implies also that we can interchange the expected value (an integral) with the infinite summation.

Weak stationarity of
$$
y_{t}
$$
.
\n**1.** $E(y_{t}) = E\left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right) = \sum_{j=0}^{\infty} \theta_{j} E(u_{t-j}) = E(u_{0}) \sum_{j=0}^{\infty} \theta_{j} = 0 \cdot \sum_{j=0}^{\infty} \theta_{j} = 0$
\n**2.** $Var(y_{t}) = E\left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right)^{2} = E\left[\left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right) \left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right)\right]$
\n
$$
= E\left[\sum_{j=0}^{\infty} \theta_{j}^{2} u_{t-j}^{2} + \sum_{j=1}^{\infty} \theta_{j} \theta_{i} u_{t-j} u_{t-i}\right] = E\left[\sum_{j=0}^{\infty} \theta_{j}^{2} u_{t-j}^{2}\right] + E\left[\sum_{j=0}^{\infty} \theta_{j} \theta_{i} u_{t-j} u_{t-i}\right]
$$

\n
$$
= \sum_{j=0}^{\infty} \theta_{j}^{2} E(u_{t-j}^{2}) + \sum_{j=1}^{\infty} \theta_{j} \theta_{i} E(u_{t-j}) E(u_{t-i}) = \sigma_{u}^{2} \sum_{j=0}^{\infty} \theta_{j}^{2} + 0 < +\infty
$$

...finite and not depending on the index.

We see that we need square summability in order to obtain a finite variance.

3. Autocovariance .

3. Autocovariance.
\n
$$
\gamma_{k} = \text{Cov}(y_{t}, y_{t-k}) = E(y_{t}y_{t-k}) - E(y_{t})E(y_{t-k}) = E(y_{t}y_{t-k})
$$
\n
$$
= E\left[\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] = E\left[\left(\sum_{j=0}^{k-1} \theta_{j}u_{t-j} + \sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right]
$$
\n
$$
= E\left[\left(\sum_{j=0}^{k-1} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right]
$$
\n
$$
= 0 + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] = E\left(\theta_{k}u_{t-k} + \theta_{k+1}u_{t-k-1} + \dots\right)\left(\theta_{0}u_{t-k} + \theta_{k+1}u_{t-k-1} + \dots\right)
$$

Cross-products are zero so we are left with
\n
$$
\gamma_k = E\Big(\theta_0 \theta_k u_{t-k}^2 + \theta_1 \theta_{k+1} u_{t-k-1}^2 + \ldots\Big) = \theta_0 \theta_k E\Big(u_{t-k}^2\Big) + \theta_1 \theta_{k+1} E\Big(u_{t-k-1}^2\Big) + \ldots
$$

Page 4 of 4
\n
$$
\Rightarrow \gamma_k = \sigma_u^2 \sum_{j=0}^{\infty} \theta_j \theta_{j+k} = \sigma_u^2 \sum_{j=0}^{\infty} \frac{\Gamma(j+d)}{\Gamma(j+1)\Gamma(d)} \frac{\Gamma(j+k+d)}{\Gamma(j+k+1)\Gamma(d)}
$$

This expression, which is just a *single* covariance term, *does not depend on the index* (it is

"summed out" in the analogous sense that we "integrate out" a variable). We can
write
 $\sigma_u^2 \sum_{i=1}^{\infty} \theta_i \theta_{j+k} = \sigma_u^2 \theta_k + \sigma_u^2 \sum_{i=1}^{\infty} \frac{1}{\Gamma(d) i^{1-d}} \frac{1}{\Gamma(d) (i+k)^{1-d}} = \sigma_u^2 \theta_k + \frac{\sigma_u^2}{(\Gamma(d))^2} \sum_{i=1}^{\infty} \frac{1}{(i+k)^{1-d}}$ write

This expression, which is just a *single* covariance term, *does not depend on the index* (it is "summed out" in the analogous sense that we "integrate out" a variable). We can write\n
$$
\gamma_k = \sigma_u^2 \sum_{j=0}^{\infty} \theta_j \theta_{j+k} = \sigma_u^2 \theta_k + \sigma_u^2 \sum_{j=1}^{\infty} \frac{1}{\Gamma(d) j^{1-d}} \frac{1}{\Gamma(d) (j+k)^{1-d}} = \sigma_u^2 \theta_k + \frac{\sigma_u^2}{(\Gamma(d))^2} \sum_{j=1}^{\infty} \frac{1}{(j^2 + jk)^{1-d}}
$$

Since *k* is fixed, the leading term in the denominator is equivalent to $\frac{1}{\sqrt{2}}$ 1 $\frac{1}{j^{2-2d}}$ which leads to convergence as long as $0 < d < 1/2$. So the autocovariance function exists for all *k*.

Also, it does not depend on the index, and certainly not on *t* , but on *k* only.

So the *y-***process is weakly stationary.**

So the *y*-process is weakly stationary.
But
$$
\sum_{k=0}^{\infty} \gamma_k > \sigma_u^2 \theta_0 + \sigma_u^2 \theta_1 + \sigma_u^2 \theta_2 + ... = \sigma_u^2 \sum_{j=0}^{\infty} \theta_j \longrightarrow \infty
$$

So the infinite sum $\boldsymbol{0}$ γ*k k* ∞ $\sum_{k=0} \gamma_k$ diverges. Therefore, the process

$$
y_{t} = \sum_{j=0}^{\infty} \theta_{j} u_{t-j}, \qquad \theta_{j} = \frac{\Gamma(j+d)}{\Gamma(j+1)\Gamma(d)}, \quad 0 < d < \frac{1}{2},
$$

is weakly stationary with "long memory".

Note: "Long memory" alone would not be a useful concept and tool. But together with weak stationarity, it enhances the concepts we have available to model and estimate real-world phenomena.

Exercises : Examine summability, weak stationarity etc in the following cases: **1.** $\mathbf{0}$ $\mu_t = \mu + \sum_{j=0} \theta_j u_{t-j},$ $y_t = \mu + \sum_{i=1}^{\infty} \theta_i u_{t-i}, \quad \mu \in \mathbb{R}$ \overline{a} $=\mu+\sum_{j=0}^{\infty}\theta_{j}u_{t-j}, \quad \mu \in \mathbb{R}$, everything else the same. **2.** For $\mathbf{0}$ $t - \sum_{j}^{\mathbf{U}} t_{j} \mathbf{u}_{t-j}$ *j* $y_i = \sum_i \theta_i u_i$ œ - $=\sum_{j=0}\theta_j u_{t-j}$ suppose that $(u_t)_{t\in\mathbb{Z}}$ has the same properties as before, except that now $E(u_t) = \delta \neq 0$.--