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An example of a "Long-memory" process

Definition: A weakly stationary stochastic process $(x_t)_{t \in \mathbb{Z}}$ with autocovariance function

 γ_k is said to have "long-memory" if $\sum_{k=0}^{\infty} \left| \gamma_k \right| = \infty$.

An example.

Consider $(u_t)_{t \in \mathbb{Z}}$ a strictly stationary White Noise process $WN(\sigma_u^2)$, and a real sequence

 $\left(heta_{j}
ight)_{j \in \mathbb{N}}$. $\forall t \in \mathbb{Z}$ we initially define

$$y_t = \sum_{j=0}^{\infty} \theta_j u_{t-j}, \qquad \theta_j = \frac{\Gamma(j+d)}{\Gamma(j+1)\Gamma(d)}, \quad d > 0,$$

where $\Gamma(\cdot)$ is the Gamma function. The parameter d is called the "memory parameter". We will see that in order for y_t to be well defined, we will have to strengthen the

restrictions on the value of *d* to 0 < d < 1/2.

The Gamma function satisfies the identity

$$\Gamma(j+1) = j\Gamma(j) \Rightarrow \theta_j = \frac{\Gamma(j+d)}{j\Gamma(j)\Gamma(d)}$$

The following exact relation holds between the Gamma function and the Beta function

$$\mathbf{B}(j,d) = \frac{\Gamma(j)\Gamma(d)}{\Gamma(j+d)} \Rightarrow \ \theta_j = \frac{1}{j\mathbf{B}(j,d)}$$

Using the Stirling approximation for "large" j and fixed/small d as is our case, we have

$$B(j,d) \approx \Gamma(d) j^{-d} \Longrightarrow \theta_j \approx \frac{1}{j \Gamma(d) j^{-d}} \Longrightarrow \theta_j \approx \frac{1}{\Gamma(d) j^{1-d}} \text{ while note that } \theta_0 = 1 \text{ exactly.}$$

Note also that $\Gamma(d)$ is bounded and does not depend on the index.

Convergence.

From the theory of infinite series we know that

$$\sum_{j=1}^{\infty} \frac{1}{j^{1-d}} \text{ diverges since } 1-d \le 1, \forall d > 0$$

So
$$\sum_{j=0}^{\infty} |\theta_j| = \sum_{j=0}^{\infty} \theta_j \longrightarrow 1 + \frac{1}{\Gamma(d)} \sum_{j=1}^{\infty} \frac{1}{j^{1-d}} \rightarrow \infty$$

And the coefficient series is not absolutely summable.

But consider

$$\sum_{j=0}^{\infty} \theta_j^2 \longrightarrow 1 + \frac{1}{\left(\Gamma(d)\right)^2} \sum_{j=1}^{\infty} \frac{1}{j^{2-2d}}$$

Whether this converges depends on the actual value of d.

If
$$0 < d < 1/2$$
 we obtain $2 - 2d > 2 - 2\frac{1}{2} = 1$, in which case $\sum_{j=0}^{\infty} \theta_j^2 < +\infty$.

Namely, for 0 < d < 1/2 we have square-summability of the coefficient series. Square-summability guarantees that $y_t = \sum_{j=0}^{\infty} \theta_j u_{t-j}$ "converges in mean-square" to some random variable (see Hamilton's book Appendix 3.A, pp 69-70).

So **we eventually impose** 0 < d < 1/2, and we have square-summability but not absolute summability. Square-summability (and hence convergence), implies also that we can interchange the expected value (an integral) with the infinite summation.

Weak stationarity of y_t .

$$1. \ E(y_{t}) = E\left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right) = \sum_{j=0}^{\infty} \theta_{j} E\left(u_{t-j}\right) = E\left(u_{0}\right) \sum_{j=0}^{\infty} \theta_{j} = 0 \cdot \sum_{j=0}^{\infty} \theta_{j} = 0$$

$$2. \ \operatorname{Var}(y_{t}) = E\left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right)^{2} = E\left[\left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right) \left(\sum_{j=0}^{\infty} \theta_{j} u_{t-j}\right)\right]$$

$$= E\left[\sum_{j=0}^{\infty} \theta_{j}^{2} u_{t-j}^{2} + \sum_{j\neq i}^{\infty} \theta_{j} \theta_{i} u_{t-j} u_{t-i}\right] = E\left[\sum_{j=0}^{\infty} \theta_{j}^{2} u_{t-j}^{2}\right] + E\left[\sum_{j\neq i}^{\infty} \theta_{j} \theta_{i} u_{t-j} u_{t-i}\right]$$

$$= \sum_{j=0}^{\infty} \theta_{j}^{2} E\left(u_{t-j}^{2}\right) + \sum_{j\neq i}^{\infty} \theta_{j} \theta_{i} E\left(u_{t-j}\right) E\left(u_{t-i}\right) = \sigma_{u}^{2} \sum_{j=0}^{\infty} \theta_{j}^{2} + 0 < +\infty$$

...finite and not depending on the index.

We see that we need square summability in order to obtain a finite variance.

3. Autocovariance .

$$\begin{split} \gamma_{k} &= \operatorname{Cov}\left(y_{t}, y_{t-k}\right) = E\left(y_{t}y_{t-k}\right) - E\left(y_{t}\right)E\left(y_{t-k}\right) = E\left(y_{t}y_{t-k}\right) \\ &= E\left[\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] = E\left[\left(\sum_{j=0}^{k-1} \theta_{j}u_{t-j} + \sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] \\ &= E\left[\left(\sum_{j=0}^{k-1} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] \\ &= 0 + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] = E\left(\theta_{k}u_{t-k} + \theta_{k+1}u_{t-k-1} + \ldots\right)\left(\theta_{0}u_{t-k} + \theta_{1}u_{t-k-1} + \ldots\right)\right) \\ &= 0 + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] = E\left(\theta_{k}u_{t-k} + \theta_{k+1}u_{t-k-1} + \ldots\right)\left(\theta_{0}u_{t-k} + \theta_{1}u_{t-k-1} + \ldots\right)\right) \\ &= 0 + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] \\ &= 0 + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-k-j}\right)\right] \\ &= 0 + E\left[\left(\sum_{j=k}^{\infty} \theta_{j}u_{t-k-j}\right)\left(\sum_{j=0}^{\infty} \theta_{j}u_{t-$$

Cross-products are zero so we are left with

$$\gamma_{k} = E\left(\theta_{0}\theta_{k}u_{t-k}^{2} + \theta_{1}\theta_{k+1}u_{t-k-1}^{2} + \dots\right) = \theta_{0}\theta_{k}E\left(u_{t-k}^{2}\right) + \theta_{1}\theta_{k+1}E\left(u_{t-k-1}^{2}\right) + \dots$$

$$\Rightarrow \gamma_k = \sigma_u^2 \sum_{j=0}^{\infty} \theta_j \theta_{j+k} = \sigma_u^2 \sum_{j=0}^{\infty} \frac{\Gamma(j+d)}{\Gamma(j+1)\Gamma(d)} \frac{\Gamma(j+k+d)}{\Gamma(j+k+1)\Gamma(d)}$$

This expression, which is just a *single* covariance term, *does not depend on the index* (it is "summed out" in the analogous sense that we "integrate out" a variable). We can write

$$\gamma_{k} = \sigma_{u}^{2} \sum_{j=0}^{\infty} \theta_{j} \theta_{j+k} = \sigma_{u}^{2} \theta_{k} + \sigma_{u}^{2} \sum_{j=1}^{\infty} \frac{1}{\Gamma(d) j^{1-d}} \frac{1}{\Gamma(d) (j+k)^{1-d}} = \sigma_{u}^{2} \theta_{k} + \frac{\sigma_{u}^{2}}{\left(\Gamma(d)\right)^{2}} \sum_{j=1}^{\infty} \frac{1}{\left(j^{2} + jk\right)^{1-d}}$$

Since *k* is fixed, the leading term in the denominator is equivalent to $\frac{1}{j^{2-2d}}$ which leads to convergence as long as 0 < d < 1/2. So the autocovariance function exists for all *k*. Also, it does not depend on the index, and certainly not on *t*, but on *k* only.

So the *y*-process is weakly stationary.

But
$$\sum_{k=0}^{\infty} \gamma_k > \sigma_u^2 \theta_0 + \sigma_u^2 \theta_1 + \sigma_u^2 \theta_2 + \dots = \sigma_u^2 \sum_{j=0}^{\infty} \theta_j \longrightarrow \infty$$

So the infinite sum $\sum_{k=0}^{\infty} \gamma_k$ diverges. Therefore, the process

$$y_t = \sum_{j=0}^{\infty} \theta_j u_{t-j}, \qquad \theta_j = \frac{\Gamma(j+d)}{\Gamma(j+1)\Gamma(d)}, \quad 0 < d < \frac{1}{2},$$

is weakly stationary with "long memory".

Note: "Long memory" alone would not be a useful concept and tool. But together with weak stationarity, it enhances the concepts we have available to model and estimate real-world phenomena.

Exercises : Examine summability, weak stationarity etc in the following cases: **1.** $y_t = \mu + \sum_{j=0}^{\infty} \theta_j u_{t-j}, \quad \mu \in \mathbb{R}$, everything else the same. **2.** For $y_t = \sum_{j=0}^{\infty} \theta_j u_{t-j}$ suppose that $(u_t)_{t \in \mathbb{Z}}$ has the same properties as before, except that now $E(u_t) = \delta \neq 0$.--