

$$\text{Cov}(\varepsilon_i, \varepsilon_j) = 0 \quad \forall i \neq j$$

AUTOCORRELATION $\begin{cases} \rightarrow \text{AR} & \text{AUTOREGRESSIVE} \\ \rightarrow \text{MA} & \text{MOVING AVERAGE} \\ \rightarrow \text{ARMA} \end{cases}$

AR(k)
 \uparrow k IS THE ORDER OF AR MODEL

$$\varepsilon_t = \mu + \phi_1 \varepsilon_{t-1} + \phi_2 \varepsilon_{t-2} + \dots + \phi_k \varepsilon_{t-k} + \eta_t$$

AR(1) $\varepsilon_t = \mu + \phi \varepsilon_{t-1} + \eta_t$

$\eta_t \stackrel{iid}{\sim} (0, \sigma^2)$

STRONG STATIONARITY

$\dots, y_{t-1}, y_t, y_{t+1}, \dots$ IS STRONGLY STATIONARY iff

$f(y_{t+1}, y_{t+2}, \dots, y_{t+k})$ is the same

$\forall t, \forall k$.

WEAK STATIONARITY OF ORDER v

$\dots, y_1, y_2, \dots, y_n, \dots$ is weakly stationary of order v iff $E(y_t^v), E(y_t^2), E(y_t^3), \dots, E(y_t^v)$ is independent of t

y_t is ~~2nd~~ 2nd ORDER STATIONARY iff
 $E(y_t), E(y_t^2), E(y_t y_{t-k})$ independent
of t .

y_t 2nd ORDER STAT. $\left. \begin{array}{l} \\ + y_t \sim N(\cdot, \cdot) \end{array} \right\} \Rightarrow y_t$ STRONGLY STATION.

y_t STRONGLY STA $\left. \begin{array}{l} \\ + V(y_t) < \infty \end{array} \right\} \Rightarrow y_t$ 2nd ORDER STATION.

AR(1)

$$\varepsilon_t = \mu + \gamma \varepsilon_{t-1} + u_t \quad u_t \stackrel{i.i.d.}{\sim} (0, \sigma^2)$$

$$E(\varepsilon_t) = \mu + \gamma E(\varepsilon_{t-1}) + \underbrace{E(u_t)}_0$$

$$\Rightarrow E(\varepsilon_t) = \mu + \gamma E(\varepsilon_{t-1})$$

if STATION., $\Rightarrow E(\varepsilon_t) = \frac{\mu}{1-\gamma}$

$$V(\varepsilon_t) = V(\mu + \gamma \varepsilon_{t-1} + u_t) =$$

$$\Rightarrow V(\varepsilon_t) = V(\gamma \varepsilon_{t-1} + u_t) \Rightarrow$$

$$\Rightarrow V(\varepsilon_t) = \gamma^2 V(\varepsilon_{t-1}) + V(u_t) + 2\gamma \text{Cov}(\varepsilon_{t-1}, u_t)$$

$$\Rightarrow V(\varepsilon_t) = \gamma^2 V(\varepsilon_{t-1}) + \sigma^2 + 0$$

IF STATION. $\Rightarrow \sqrt{\text{Var}(\varepsilon_t)} = \frac{\sigma^2}{1-\gamma^2}$

$$\varepsilon_t = \mu + \gamma \varepsilon_{t-1} + u_t =$$

$$= \mu + \gamma \mu + \gamma^2 \varepsilon_{t-2} + u_t + \gamma u_{t-1},$$

$$= \mu + \gamma \mu + \gamma^2 \mu + \gamma^3 \varepsilon_{t-3} + u_t + \gamma u_{t-1} + \gamma^2 \varepsilon_{t-2}$$

$$= \mu(1 + \gamma + \gamma^2 + \dots) + u_t + \gamma u_{t-1} + \gamma^2 u_{t-2} + \gamma^3 u_{t-3} + \dots$$

$$\Rightarrow E(\varepsilon_t) = \mu(1 + \gamma + \gamma^2 + \gamma^3 + \dots)$$

$$\Rightarrow E(\varepsilon_t) = \frac{\mu}{1-\gamma} \quad \text{iff } |\gamma| < 1 \quad \textcircled{4}$$

$$V(z_t) = V(n(1 + \gamma + \gamma^2 + \dots) + n_t + \gamma n_{t-1} + \gamma^2 n_{t-2} + \dots)$$

$$= V(n_t + \gamma n_{t-1} + \gamma^2 n_{t-2} + \gamma^3 n_{t-3} + \dots)$$

$$= V(n_t) + \gamma^2 V(n_{t-1}) + \gamma^4 V(n_{t-2}) + \gamma^6 V(n_{t-3}) + \dots$$

$$\Rightarrow V(z_t) = \sigma^2 (1 + \gamma^2 + \gamma^4 + \gamma^6 + \dots)$$

$$\Rightarrow V(z_t) = \frac{\sigma^2}{1 - \gamma^2} \stackrel{\text{IA}}{=} \gamma^2 < 1 \quad \textcircled{B}$$

$$\textcircled{A} + \textcircled{B} \Rightarrow | \gamma | < 1$$

$$\gamma_k = \text{Cov}(\varepsilon_t, \varepsilon_{t-k})$$

$$\gamma_1 = \text{Cov}(\varepsilon_t, \varepsilon_{t-1})$$

$$\text{Cov}(\varepsilon_{t+1}, \varepsilon_t) = \gamma_1$$

$$\varepsilon_t = \mu + \gamma \varepsilon_{t-1} + u_t \rightarrow$$

$$\text{Cov}(\varepsilon_t, \varepsilon_{t-k}) = \text{Cov}(\mu + \gamma \varepsilon_{t-1} + u_t, \varepsilon_{t-k})$$

$$\Rightarrow \gamma_k = \gamma \text{Cov}(\varepsilon_{t-1}, \varepsilon_{t-k}) + \text{Cov}(u_t, \varepsilon_{t-k})$$

$$\Rightarrow \gamma_k = \gamma \gamma_{k-1} + 0$$

$$\Rightarrow r_k = \rho r_{k-1}$$

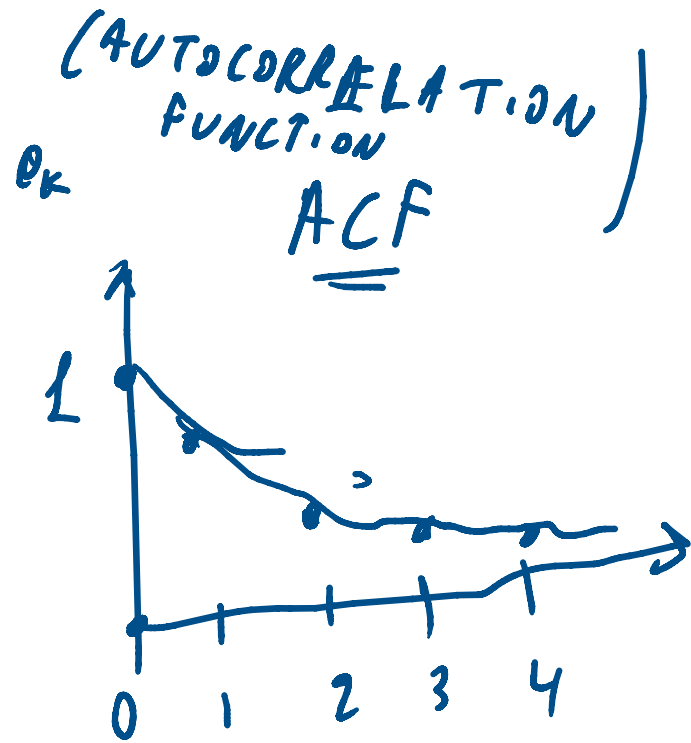
$$r_1 = \rho r_0 = \rho \frac{\sigma^2}{1-\rho^2}$$

$$r_2 = \rho r_1 = \rho^2 r_0$$

\vdots

$$r_k = \rho^k r_0$$

$$\rho_k = \frac{r_k}{r_0} = \rho^k \leftarrow$$



ρ_k° = PARTIAL AUTOCOR. OF ORDER k

= ΑΥΤΟCΟΛ. ΤΟΥ ε_t ΜΕ ε_{t-k} ΑΦΟΪ
ΛΑΪΒΩ ΥΠ' ΟΨΗΝ ΜΟΝ ΤΑ ΑΥΤΟCΟΛ. ΤΟΥ
($\varepsilon_t, \varepsilon_{t-1}$), ($\varepsilon_t, \varepsilon_{t-2}$) ..., ($\varepsilon_t, \varepsilon_{t-k+1}$)

$$\varepsilon_t = c^* + \rho_k \varepsilon_{t-k} + v_t$$

$$\varepsilon_t = c^{*0} + \sigma_1 \varepsilon_{t-1} + \sigma_2 \varepsilon_{t-2} + \dots + \sigma_{k-1} \varepsilon_{t-k+1} + \rho_k^\circ \varepsilon_{t-k} + v_t^\circ$$

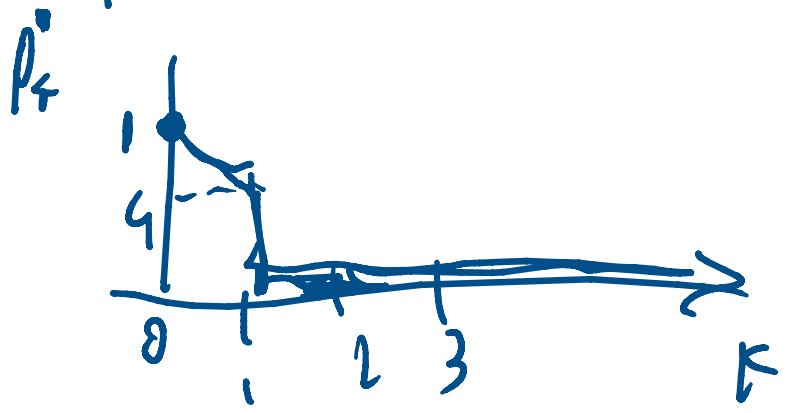
$$\rho_1^* = \rho_1 = 4$$

$$\varepsilon_t = \mu + \gamma \varepsilon_{t-1} + \rho_2^* \varepsilon_{t-2} + u_{t-1}$$

ε_t öböl $E \varepsilon_t = 0$ $\varepsilon_t = \mu + \gamma \varepsilon_{t-1} + u_{t-1}$

\Rightarrow

$$\Rightarrow \rho_2^* = 0 \sim \rho_3^* \sim \rho_4^* = \dots$$



$$y_t = x_t \beta + \varepsilon_t$$

1×1 $1 \times k$ $k \times 1$

$$\varepsilon_t = \rho \varepsilon_{t-1} + u_t, \quad u_t \stackrel{iid}{\sim} (0, \sigma^2)$$

$$Y = X\beta + \varepsilon \quad E(\varepsilon) = 0$$

$n \times 1$ $n \times k$ $k \times 1$ $n \times 1$

$$V(\varepsilon) = E(\varepsilon \varepsilon')$$

$$= E \begin{pmatrix} \varepsilon_1^2 & \varepsilon_1 \varepsilon_2 & \varepsilon_1 \varepsilon_3 & \dots & \varepsilon_1 \varepsilon_n \\ \varepsilon_2 \varepsilon_1 & \varepsilon_2^2 & \rho \varepsilon_2 \varepsilon_3 & \dots & \rho \varepsilon_2 \varepsilon_n \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \varepsilon_n \varepsilon_1 & \varepsilon_n \varepsilon_2 & \dots & \dots & \varepsilon_n^2 \end{pmatrix} = \begin{pmatrix} \sigma_0 & \sigma_1 & \sigma_2 & \dots & \sigma_{n-1} \\ \sigma_1 & \sigma_0 & \sigma_1 & \dots & \sigma_{n-2} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \sigma_{n-1} & \dots & \dots & \dots & \sigma_0 \end{pmatrix} =$$

$$= \sigma_0 \begin{pmatrix} 1 & \rho_1 & \rho_2 & \dots & \rho_{n-1} \\ \rho_1 & 1 & \rho_1 & & \rho_{n-2} \\ \vdots & & \vdots & & \vdots \\ \rho_{n-1} & \rho_{n-1} & \dots & & 1 \end{pmatrix} =$$

$$= \frac{\sigma^2}{1-\rho^2} \begin{pmatrix} 1 & \rho & \rho^2 & \dots & \rho^{n-1} \\ \rho & 1 & \rho & \dots & \rho^{n-2} \\ \vdots & & \vdots & & \vdots \\ \rho^{n-1} & \rho^{n-1} & \dots & & 1 \end{pmatrix} = \sigma^2 \underline{Q}$$

$V(\hat{\beta}) \neq \sigma^2 I_n \Rightarrow \hat{\beta} = (X'X)^{-1}X'y$ is UNBIASED
BUT INEFFICIENT.

$$\varepsilon_t = \mu + u_t - \theta u_{t-1}, \quad \text{MA}(1)$$

$$E(\varepsilon_t) = \mu + \theta - \theta \cdot 0 = \mu \quad \forall \theta$$

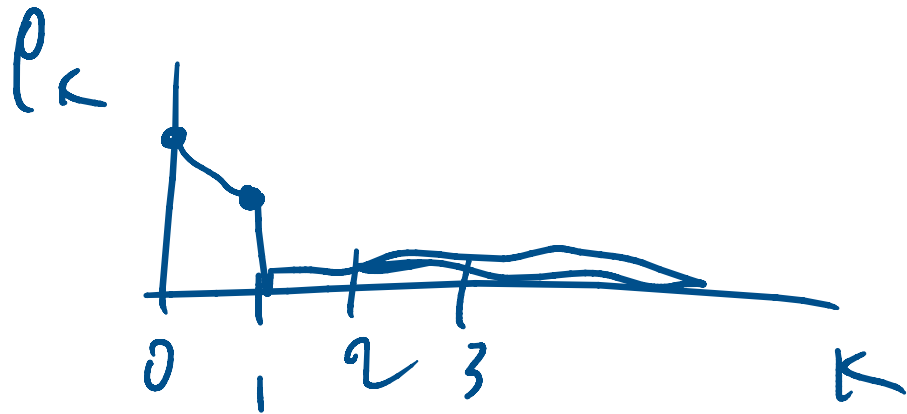
$$V(\varepsilon_t) = V(\mu + u_t - \theta u_{t-1}) = V(u_t) + \theta^2 V(u_{t-1}) - 2\theta \text{Cov}(u_t, u_{t-1}) = \sigma^2 + \sigma^2 \theta^2 - 0 =$$

$$\Rightarrow V(\varepsilon_t) = \sigma^2 (1 + \theta^2) \quad \forall \theta,$$

\Rightarrow MA(1) ΠΑΝΤΑ ΣΤΑΘΙΜΟ ($\forall \theta$ is stationary 2^{nd} order)

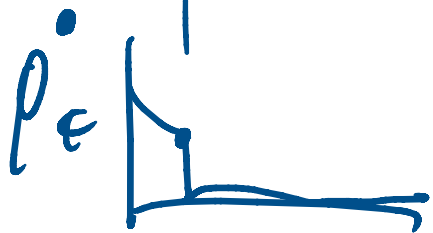
$$\text{Cov}(\varepsilon_t, \varepsilon_{t-k}) = \gamma_k = \begin{cases} \frac{-\theta}{(1+\theta^2)} \sigma^2 & k=1 \\ 0 & \forall k > 1 \end{cases}$$

$$\rho_k = \begin{cases} -\frac{\vartheta}{1+\vartheta^2} & k=1 \\ 0 & k \geq 2 \end{cases}$$

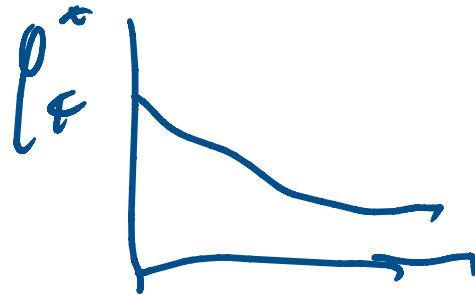
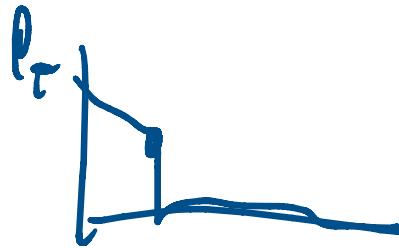


$$\rho_k^* = \vartheta^k \cdot \text{ct.}$$

AR(1)



MA(1)



ARMA(1,1)

$$\varepsilon_t = \mu + \phi\varepsilon_{t-1} + \eta_t - \theta\varepsilon_{t-1}$$



AR(k), MA(k), ARMA(p, q)

ESTIMATION VIA MAX. LIKELIHOOD...

$$Y = X\beta + \varepsilon \quad E(\varepsilon) = 0, \quad V(\varepsilon) = \sigma^2 I_n$$

$n \times 1$ $n \times k$ $k \times 1$ $n \times 1$

$\hat{\beta} = (X'X)^{-1} X'Y$ is the estimator
that minimizes $\sum_{t=1}^n \varepsilon_t^2 = \varepsilon' \varepsilon$

$$\varepsilon \sim N(0, \sigma^2 I_n)$$

$$\Rightarrow L(\beta_1, \beta_2, \dots, \beta_n),$$

$$\max_{\beta, \sigma^2} L(y_1, y_2, \dots, y_n) \rightarrow \hat{\beta}_{MLE} = \hat{\beta}$$

$\hat{\sigma}^2 = \frac{1}{n} \sum \varepsilon_t^2$

$$X \sim \chi^2_n$$

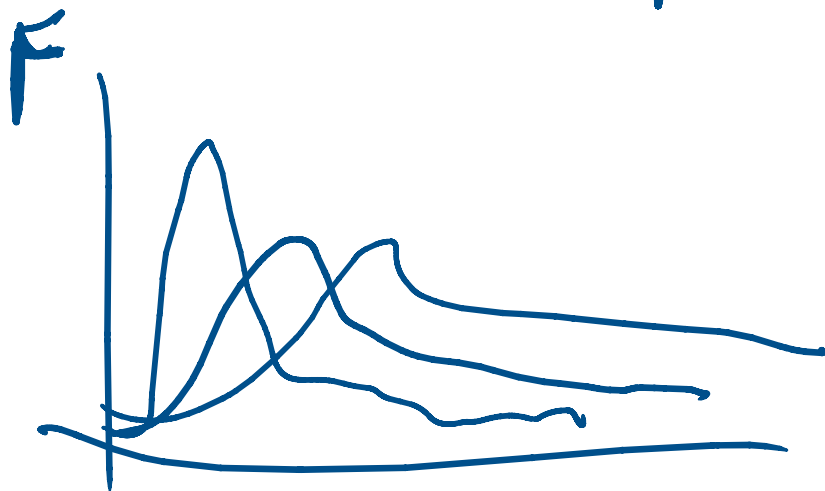
$$Y \sim \chi^2_k$$

X, Y INDEP.



\Rightarrow

$$\frac{X/n}{Y/k} \sim F(n, k)$$



$$\rho_k = \text{corr}(y_t, y_{t-k})$$

$$H_0: \rho_k = 0 \text{ v.s. } H_1: \rho_k \neq 0$$

$$\sqrt{n} \hat{\rho}_k \stackrel{H_0}{\underset{A}{\rightsquigarrow}} N(0, 1) \Rightarrow$$

$$\Rightarrow n \hat{\rho}_k^2 \stackrel{H_0}{\underset{A}{\rightsquigarrow}} \chi^2_1$$

$$\text{cov}(\hat{\rho}_k, \hat{\rho}_\lambda) = 0 \quad \forall k \neq \lambda$$

TESTING AUTOCOR.

$$H_0: \rho_k = 0 \text{ v.s. } H_1: \rho_k \neq 0$$

$$\hat{\rho}_k \underset{H_0}{\overset{A}{\sim}} N\left(0, \frac{1}{n}\right) \Leftrightarrow \sqrt{n} \hat{\rho}_k \underset{H_0}{\overset{A}{\sim}} N(0, 1)$$

Reject H_0 iff

$$|\sqrt{n} \hat{\rho}_k| > 1.96 \quad (5\% \text{ size 2-sided})$$

$$\Rightarrow \sqrt{n} \hat{\rho}_k \underset{H_0}{\overset{A}{\sim}} \chi^2_1 \Rightarrow \text{Reject } H_0 \text{ iff } \sqrt{n} \hat{\rho}_k^2 > \chi^2_{1, 5\%}$$

$$Q_k = n(\rho_1^2 + \rho_2^2 + \dots + \rho_k^2) \underset{A}{\overset{H_0}{\sim}} \chi_k^2$$

$$H_0: \rho_1 = \rho_2 = \dots = \rho_k = 0 \quad \text{v.s.} \quad H_1:$$

At least one $\neq 0$

$$y_t = x_t' \hat{\beta} + \hat{\varepsilon}_t$$

$$\hat{\varepsilon}_t = x_t' \hat{\delta} + \alpha_1 \hat{\varepsilon}_{t-1} + \alpha_2 \hat{\varepsilon}_{t-2} + \dots + \alpha_k \hat{\varepsilon}_{t-k}$$

$$\downarrow R^2$$

$$nR^2$$

$$\underset{A}{\overset{H_0}{\sim}} \chi_k^2$$

$$H_0: \alpha_1 = \alpha_2 = \dots = \alpha_k = 0$$

$$y_t = x_t' \beta + \varepsilon_t$$

$$\varepsilon_t = \eta_t - \rho \eta_{t-1} \quad \eta_t \stackrel{iid}{\sim} (0, \sigma^2)$$

MOVING AVERAGE of order 1 \equiv MA(1)
 TO ESTIMATE ρ

PORTMANTEAU Q_k
 BRUSH - GODFREY LM

